

Task-Driven Uncertainty Quantification in Inverse Problems via Conformal Prediction

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Introduction

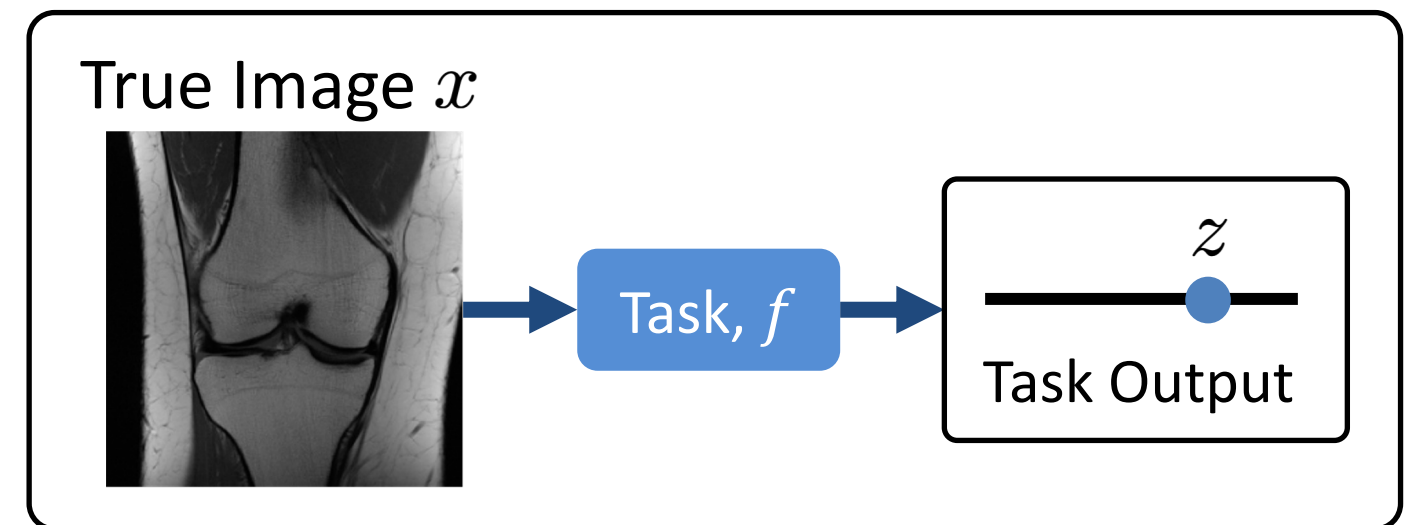
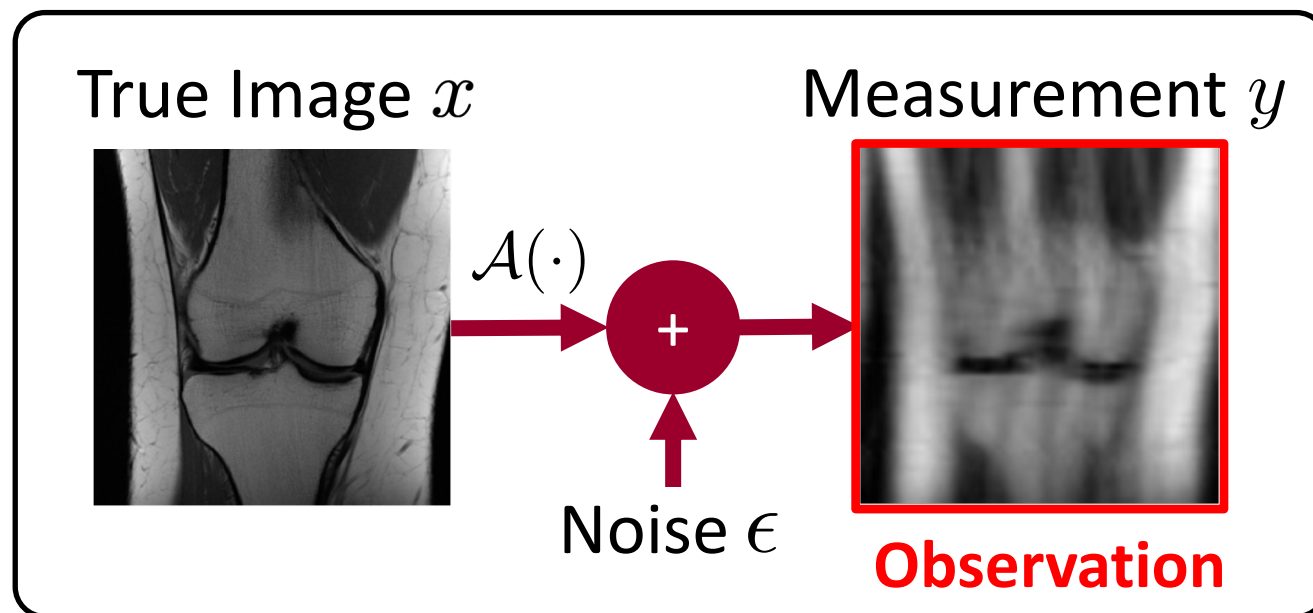
➤ Consider

▪ Inverse problem $y = \mathcal{A}(x) + \epsilon$

- Accelerated MRI
- Inpainting, super-resolution, etc

▪ Imaging task $f(x) \in \mathbb{R}$

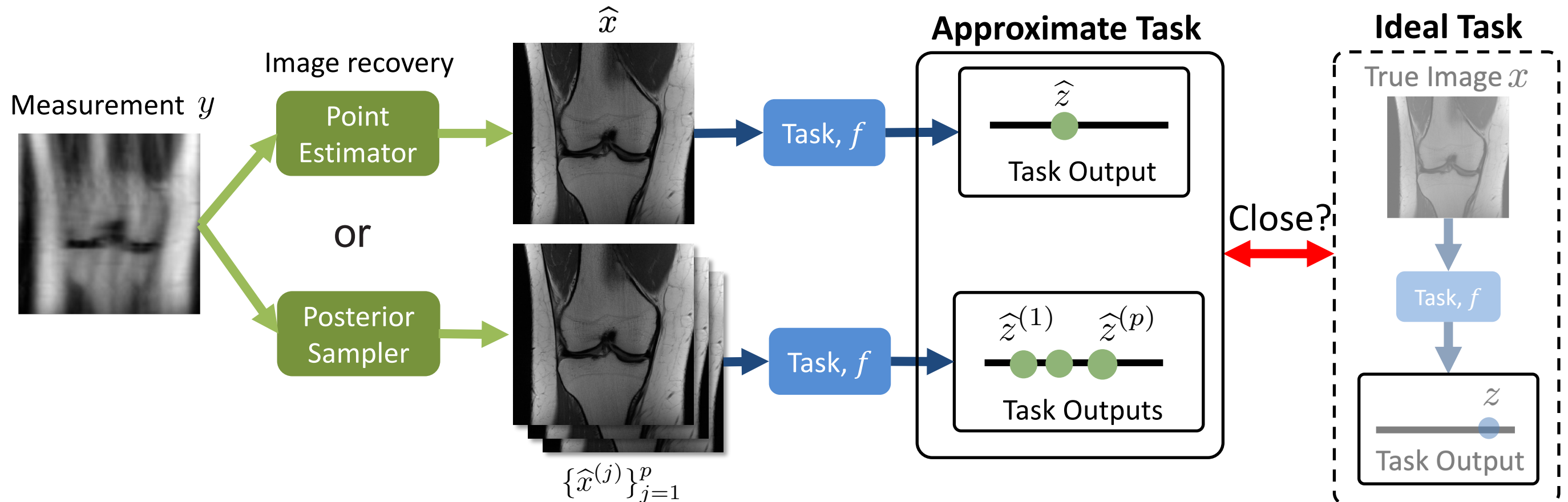
- Classification (soft-output)
- Tumor size estimation



➤ Goal: Perform the task given only the measurements y

Introduction

- Could approximate the task using reconstructed image(s) $\{\hat{x}^{(j)}\}_{j=1}^p$

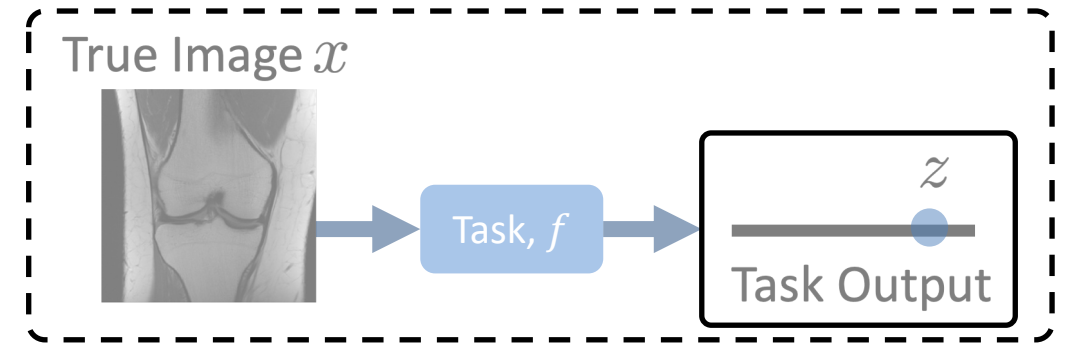


- Rather than quantifying the accuracy or uncertainty of $\hat{z}$ or $\hat{z}^{(j)}$,
we **compute rigorous bounds on the true value of z itself**

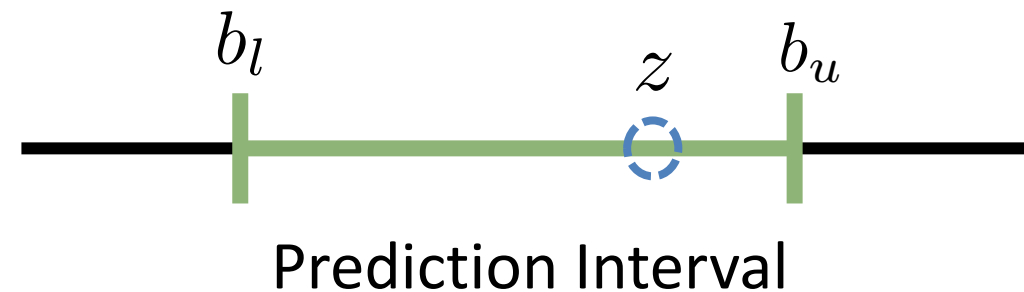
Objective

➤ Given:

- Measurement y
- Image reconstruction method/sampler $y \mapsto \{\hat{x}^{(j)}\}$
- Calibration set



➤ Goal: Construct prediction interval $\mathcal{C}_{\hat{\lambda}}(\{\hat{x}^{(j)}\}) = [b_l, b_u]$ guaranteed to contain the unknown true task output z with high probability



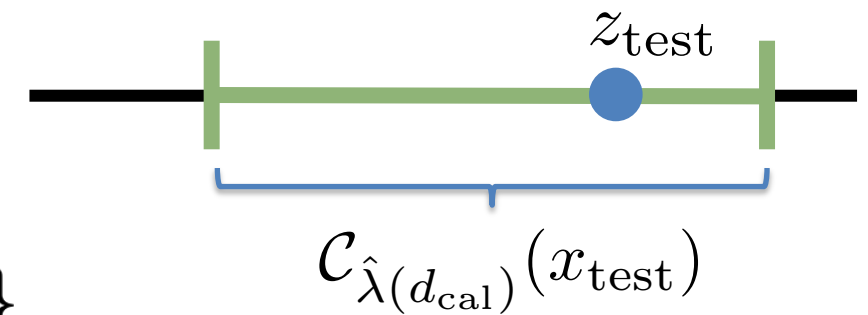
Where: $\mathbb{P}(Z \in \mathcal{C}_{\hat{\lambda}(D_{\text{cal}})}(\{\hat{X}^{(j)}\})) \geq 1 - \boxed{\alpha}$

User-chosen error rate

Main tool: conformal prediction

➤ Setup

- Given a black-box predictor $f : \mathcal{X} \rightarrow \mathcal{Z}$
- Given a calibration set $d_{\text{cal}} = \{(x_i, z_i)\}_{i=1}^n$
- Choose an error rate α
- Construct a prediction set $\mathcal{C}_\lambda(x) \subset \{\text{subsets of } \mathcal{Z}\}$



- Compute $\hat{\lambda}(d_{\text{cal}})$ such that $\mathbb{P}(Z_{\text{test}} \in \underbrace{\mathcal{C}_{\hat{\lambda}(D_{\text{cal}})}(X_{\text{test}})}_{\text{on expectation over random calibration set and test sample}}) \geq 1 - \alpha$

Holds when $\{(X_1, Z_1), \dots, (X_n, Z_n), (X_{\text{test}}, Z_{\text{test}})\}$ are exchangeable

Constructing the prediction set

1. Choose a nonconformity score $s(x, z; f) \in \mathbb{R}$
 - Higher values of s correspond to worse predictions of z

2. Compute scores for calibration samples:

$$s_i = s(x_i, z_i; f), \quad i = 1, \dots, n$$

3. Compute empirical quantile

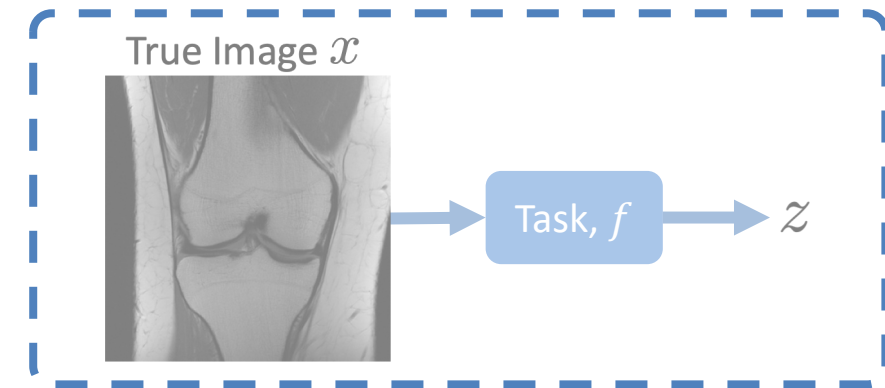
$$\hat{\lambda}(d_{\text{cal}}) = \text{Quantile}\left(\frac{\lceil (1-\alpha)(n+1) \rceil}{n}; s_1, \dots, s_n\right)$$

4. Construct prediction set

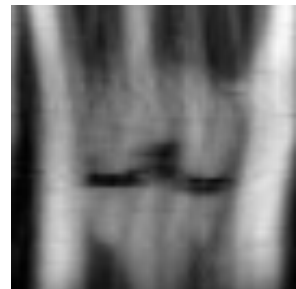
$$\mathcal{C}_{\hat{\lambda}(d_{\text{cal}})}(x_{\text{test}}) = \{z : s(x_{\text{test}}, z; f) \leq \hat{\lambda}(d_{\text{cal}})\}$$

Our method: Overview

- Goal: Bound the true task output z



Measurements y



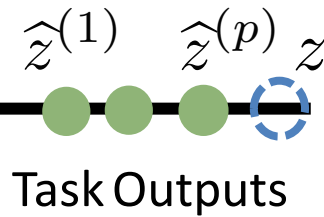
Reconstruction Method



Recovered Image(s),
 $\{\hat{x}^{(j)}\}_{j=1}^p$

Task, f

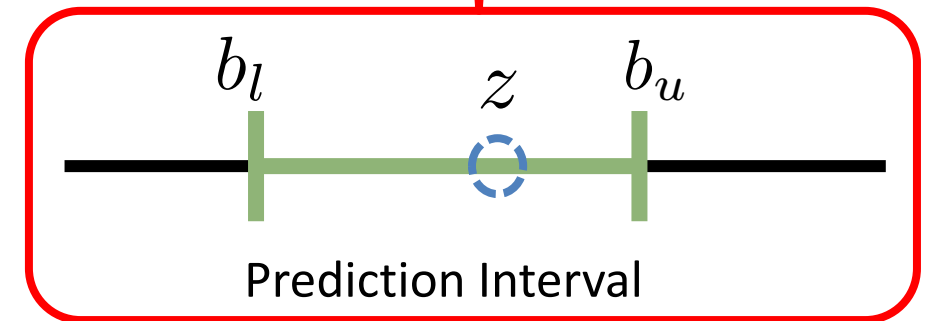
$\{\hat{z}^{(j)}\}_{j=1}^p$



$$d_{\text{cal}} = \{(\{\hat{z}_i^{(j)}\}_{j=1}^p, z_i)\}_{i=1}^n$$

Conformal Prediction

$$\mathcal{C}_{\hat{\lambda}(d_{\text{cal}})}(\{\hat{x}^{(j)}\}) = [b_l, b_u]$$

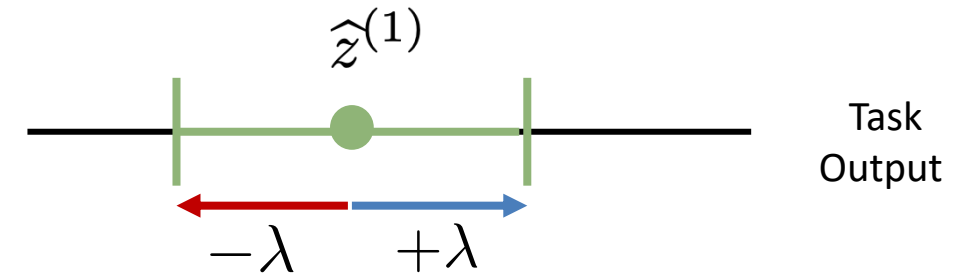


$$\text{Where: } \mathbb{P}(Z \in \mathcal{C}_{\hat{\lambda}(D_{\text{cal}})}(\{\hat{X}^{(j)}\})) \geq 1 - \alpha$$

Our method: Three ways to construct prediction sets

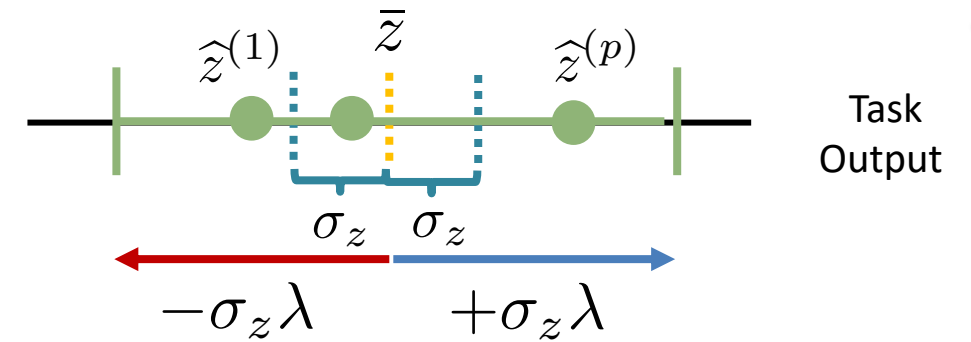
➤ Absolute Residual (AR)

$$C_\lambda(\hat{x}^{(1)}) = [\hat{z}^{(1)} - \lambda, \hat{z}^{(1)} + \lambda]$$



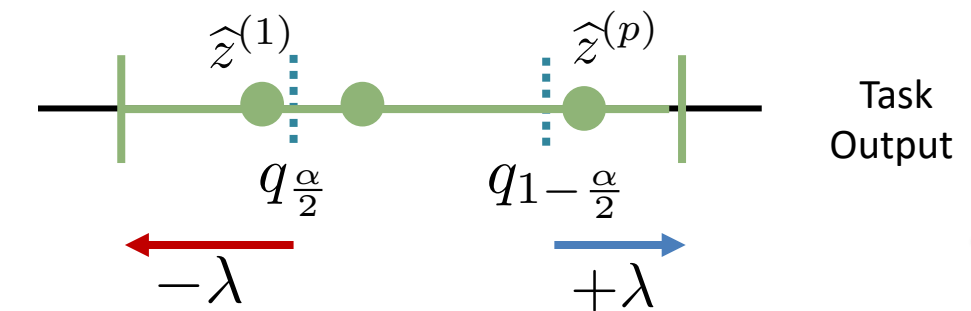
➤ Locally-Weighted Residual (LWR)

$$C_\lambda(\{\hat{x}^{(j)}\}) = [\bar{z} - \sigma_z \lambda, \bar{z} + \sigma_z \lambda] \begin{cases} \bar{z} = \frac{1}{p} \sum_{j=1}^p \hat{z}^{(j)} \\ \sigma_z^2 = \frac{1}{p} \sum_{j=1}^p (\hat{z}^{(j)} - \bar{z})^2 \end{cases}$$



➤ Conformalized Quantile Regression (CQR)

$$C_\lambda(\{\hat{x}^{(j)}\}) = [q_{\frac{\alpha}{2}}(\{\hat{z}^{(j)}\}) - \lambda, q_{1-\frac{\alpha}{2}}(\{\hat{z}^{(j)}\}) + \lambda]$$



Adaptive $|C_\lambda|$

Experiments

- Dataset: fastMRI multicoil knee, non-fat-suppressed
- Task: Soft-output meniscus-tear classification
- Reconstruction models
 - Non-adaptive: E2E-VarNet point estimator
 - Adaptive: Conditional Normalizing Flow (CNF) posterior sampler

Zbontar et al., fastMRI: An open dataset and benchmarks for accelerated MRI. *arXiv:1811.08839*, 2018.

Sriram, Zbontar, Murrell, Defazio, Zitnick, Yakubova, Knoll, and Johnson. End-to-end variational networks for accelerated MRI reconstruction. *Proc. MICCAI*, 2020.

Wen, Ahmad, and Schniter. A conditional normalizing flow for accelerated multi-coil MR imaging. *Proc. ICML*, 2023.

Empirical verification of coverage $\mathbb{P}(Z_{\text{test}} \in C_{\hat{\lambda}(D_{\text{cal}})}) \geq 1 - \alpha$

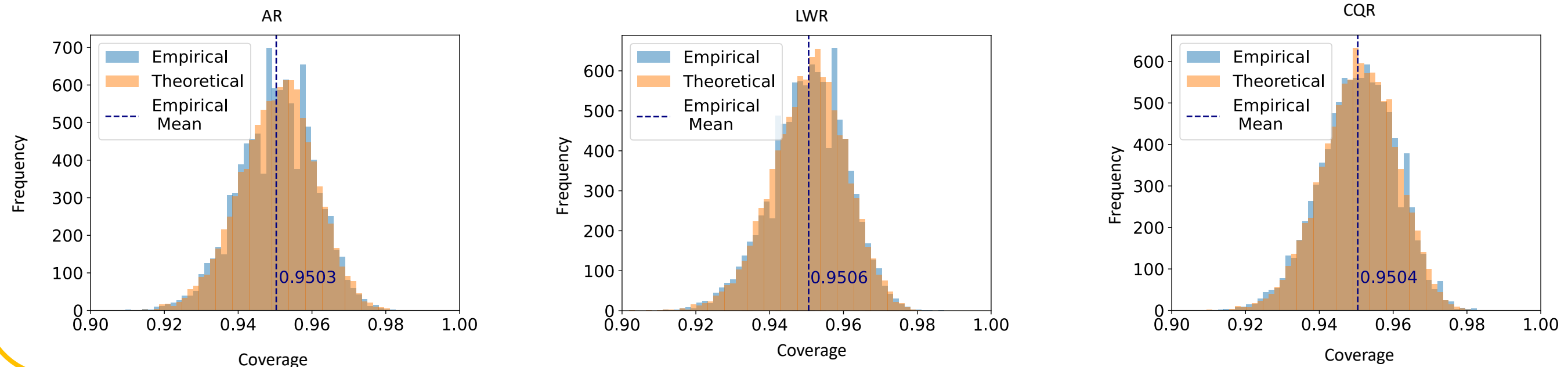
➤ Perform $T = 10000$ Monte Carlo trials:

- Randomly split validation set

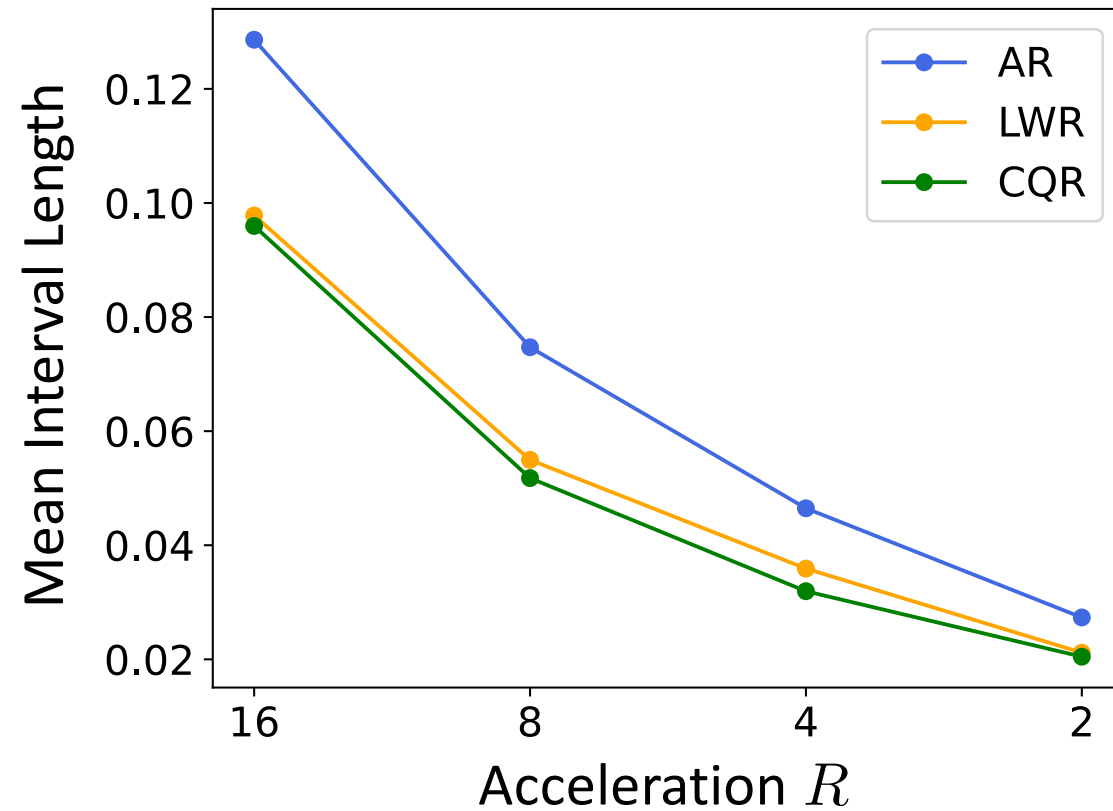
- 70% for calibration
- 30% for testing

- Compute empirical coverage
$$\text{EC}[t] = \frac{1}{|\mathcal{I}_{\text{test}}[t]|} \sum_{i \in \mathcal{I}_{\text{test}}[t]} \mathbb{1} \left\{ z_i \in C_{\hat{\lambda}(d_{\text{cal}}[t])}(\{\hat{x}_i^{(j)}\}) \right\}$$

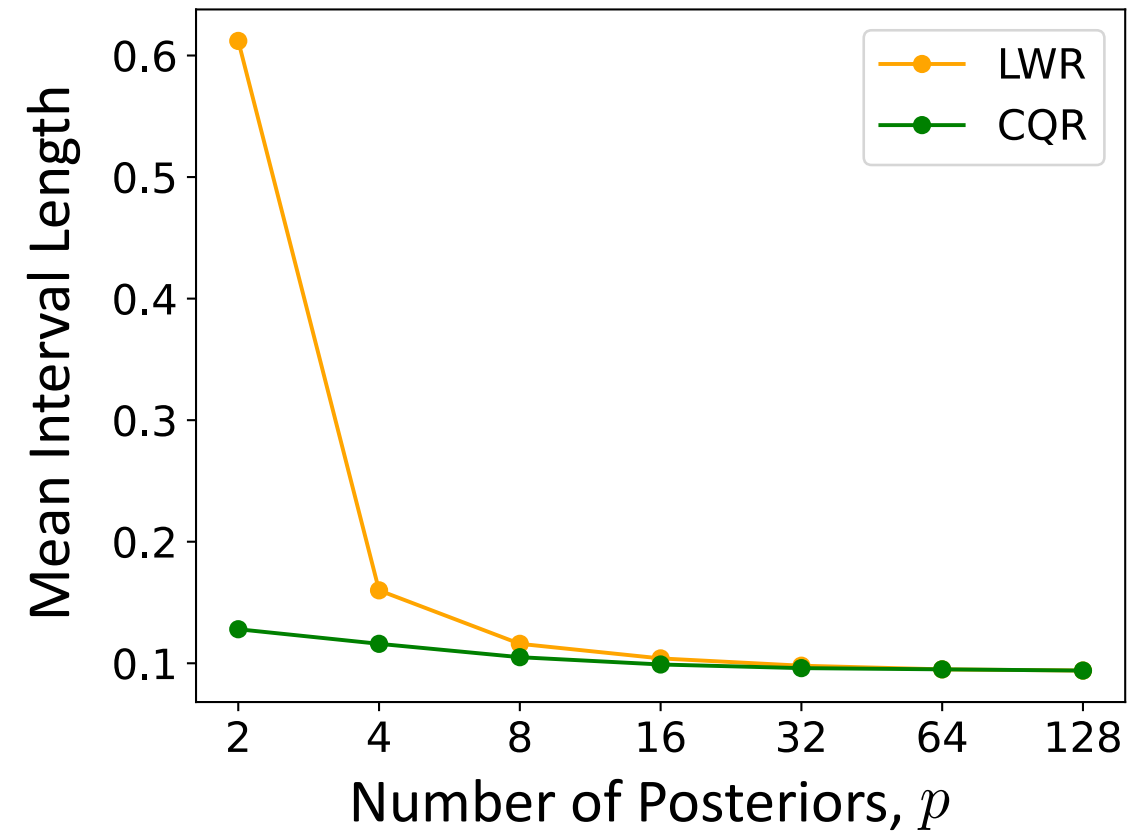
Histograms of Empirical Coverage for $\alpha = 0.05$



Comparison of prediction methods



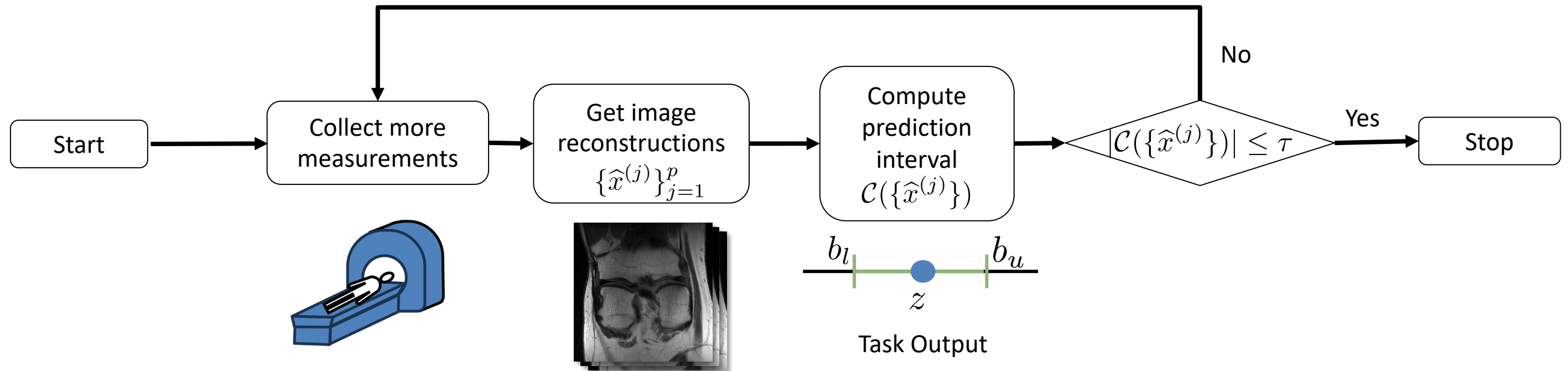
(a) Mean Interval Length vs. R



(b) Mean Interval Length vs p

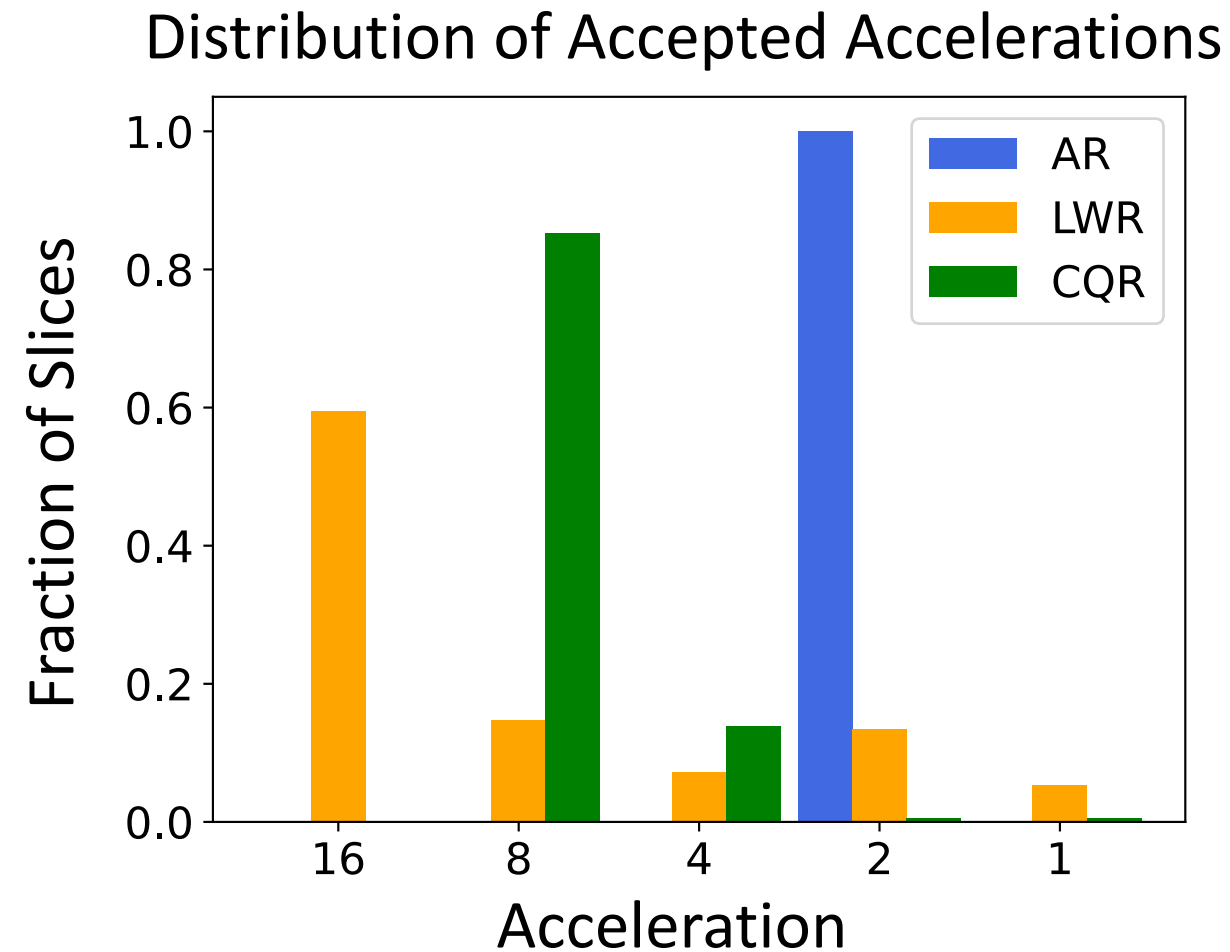
- Adaptive schemes yield smaller uncertainty intervals $|C_{\hat{\lambda}(d_{\text{cal}})}|$
- CQR is robust to the use of very few posterior samples p

A multi-round measurement scheme with uncertainty guarantees



Quantitative results for the multi-round measurement scheme

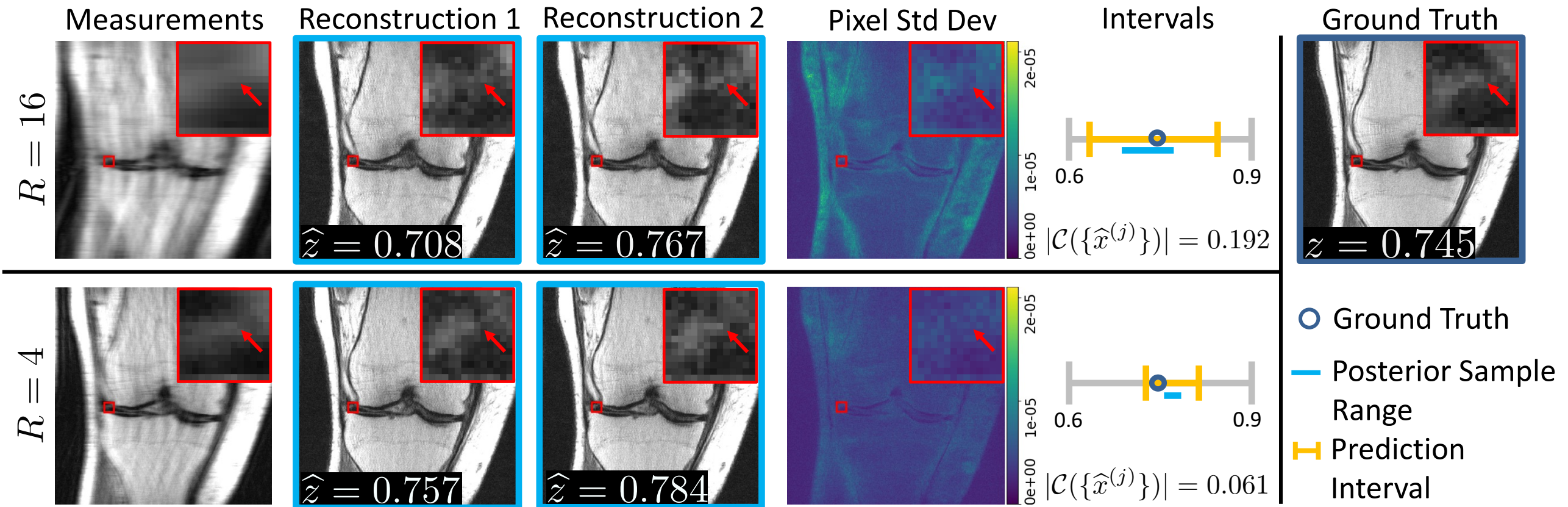
- Error rate $\alpha = 0.01$
- Threshold $\tau = 0.1$



Method	Average Acceleration	Empirical Coverage
AR	2.000	0.991 ± 0.008
LWR	5.157	0.992 ± 0.005
CQR	6.762	0.987 ± 0.008

$\pm$ Standard Error

Visual examples for the multi-round measurement scheme



➤ If the threshold was $\tau = 0.1$ then $R = 4$ would suffice, but not $R = 16$

Conclusion

- Goal: Construct an interval guaranteed to contain the true task output z with probability $\geq 1 - \alpha$
- Accomplished this using conformal prediction
 - Investigated three prediction methods: AR, LWR, CQR
 - LWR and CQR leverage a posterior sampler to adapt $|C_{\hat{\lambda}}|$ to y
 - CQR is robust to very few posterior samples
- Proposed a multi-round protocol with uncertainty guarantees
 - Adaptive schemes offer large acceleration gains over non-adaptive

Intuition behind conformal prediction

- Suppose $D_{\text{cal}} = \{Z_i\}_{i=1}^n$ with i.i.d $\{Z_1, \dots, Z_n, Z_{\text{test}}\}$
 - We're ignoring the feature X to keep things simple

- Suppose we want to find $C_\lambda = (-\infty, \lambda]$ and $\hat{\lambda}(D_{\text{cal}})$ such that

$$\mathbb{P}(Z_{\text{test}} \in C_{\hat{\lambda}(D_{\text{cal}})}) \geq 1 - \alpha$$

- If we set $\hat{\lambda}(D_{\text{cal}}) = \text{Quantile}(1 - \alpha; Z_1, \dots, Z_n)$ then

$$\mathbb{P}(Z_{\text{test}} \leq \hat{\lambda}(D_{\text{cal}})) = \mathbb{P}(Z_{\text{test}} \in C_{\hat{\lambda}(D_{\text{cal}})}) = 1 - \alpha \quad \text{as } n \rightarrow \infty$$

- Can we do something similar that holds with finite n ?

Intuition behind conformal prediction (cont.)

➤ Note that, since $\{Z_1, \dots, Z_n, Z_{\text{test}}\}$ are i.i.d, the rank of Z_{test} is uniform in $\{1, 2, \dots, n + 1\}$

➤ This implies that

$$\begin{aligned} & \mathbb{P}(Z_{\text{test}} \leq \text{the } \lceil (1 - \alpha)(n + 1) \rceil \text{ smallest of } Z_1, \dots, Z_n, Z_{\text{test}}) \geq 1 - \alpha \\ \Leftrightarrow & \mathbb{P}(Z_{\text{test}} > \text{the } \lceil (1 - \alpha)(n + 1) \rceil \text{ smallest of } Z_1, \dots, Z_n, Z_{\text{test}}) < \alpha \\ \Leftrightarrow & \mathbb{P}(Z_{\text{test}} > \text{the } \lceil (1 - \alpha)(n + 1) \rceil \text{ smallest of } Z_1, \dots, Z_n) < \alpha \\ \Leftrightarrow & \mathbb{P}(Z_{\text{test}} \leq \text{the } \lceil (1 - \alpha)(n + 1) \rceil \text{ smallest of } Z_1, \dots, Z_n) \geq 1 - \alpha \end{aligned}$$

➤ Equivalently,

$$\mathbb{P}(Z_{\text{test}} \leq \hat{\lambda}(D_{\text{cal}})) = \mathbb{P}(Z_{\text{test}} \in C_{\hat{\lambda}(D_{\text{cal}})}) \geq 1 - \alpha$$

$$\text{when } \hat{\lambda}(D_{\text{cal}}) = \text{Quantile}\left(\frac{\lceil (1 - \alpha)(n + 1) \rceil}{n}; Z_1, \dots, Z_n\right)$$