Non-Random Parameter Estimation

Given an observation y drawn from $p_{\theta}(y)$, how do we estimate the parameter θ ? Here we assume **no** prior statistical knowledge of θ .

- 1. Terminology/Formulae:
 - (a) <u>Bias</u>

The bias of estimator $\hat{\theta}$ is $b(\hat{\theta}) = \mathbf{E}_{\theta} \left\{ \hat{\theta}(y) \right\} - \theta.$

An estimator is *unbiased* if $b(\hat{\theta}) = 0$.

(b) <u>Sufficient statistic</u>

A function T(y) is a sufficient statistic for estimation of θ if $P_{\theta}(y|T(y))$ is not dependent on θ .

(c) Minimal sufficient statistic

A sufficient statistic for θ is minimal if it is a function of every other sufficient statistic for θ .

(d) Jensen's Inequality

If f(x) is a convex function, then $\mathbf{E}{f(X)} \ge f(\mathbf{E}{X})$ with equality iff $X = E{X}$ w.p. 1.

(e) Neyman-Fisher Factorization Theorem T(y) is a sufficient statistic for θ iff there exist functions g_{θ} and h(y) such that

$$p_{\theta}(y) = g_{\theta}(T(y))h(y) \ \forall \ \theta \in \Lambda, y \in \Gamma$$

(f) Rao-Blackwell Theorem

Suppose that $\hat{g}(y)$ is any unbiased estimator of $g(\theta)$ and that T(y) is a sufficient statistic for θ . Define

$$\tilde{g}(T(y)) = \mathbf{E}_{\theta} \left\{ \hat{g}(Y) | T(Y) = T(y) \right\},\$$

Then

- $\tilde{g}(T(y))$ is an unbiased estimator of $g(\theta)$
- $\operatorname{Var}_{\theta} \{ \tilde{g}(Y) \} \leq \operatorname{Var}_{\theta} \{ \hat{g}(Y) \}$ with equality iff $P_{\theta} [\tilde{g}(Y) = \hat{g}(Y)] = 1$.
- (g) Complete family

The family $\{P_{\theta}; \theta \in \Lambda\}$ is said to be complete when $\mathbf{E}_{\theta} \{f(Y)\} = 0 \quad \forall \ \theta \in \Lambda \Rightarrow P_{\theta} [f(Y) = 0] = 1$. Note that if $\{P_{\theta}; \theta \in \Lambda\}$ is complete, all sufficient statistics for θ are trivial, i.e., every sufficient statistic is related to the observation by a 1:1 function.

(h) Complete sufficient statistic

Suppose T(y) is sufficient for $\{P_{\theta}; \theta \in \Lambda\}$ and $T(Y) \sim Q_{\theta}$ when $Y \sim P_{\theta}$.

If $\{Q_{\theta}; \theta \in \Lambda\}$ is a complete family, then T(y) is a complete sufficient statistic.

Any unbiased estimator that is a function of a complete sufficient statistic is unique and is the MVUE! Complete sufficient statistics are minimal.

(i) Exponential Family

A class of distributions is said to be an exponential family if there exist the real-valued functions $C, Q_1, \dots, Q_m, T_1, \dots, T_m$, and h such that

$$p_{\theta}(y) = C(\theta)h(y)\exp\left\{\sum_{l=1}^{m} Q_{l}(\theta)T_{l}(y)\right\} \quad \forall y \in \Gamma \subset \mathbb{R}^{n}, \ \forall \theta \in \Lambda \subset \mathbb{R}^{m}$$

The T_l above are complete sufficient statistics when Λ contains an *m*-dimensional rectangle.

(j) The Information Inequality

Under suitable regularity conditions,

$$\begin{aligned} \mathbf{Var}_{\theta} \left[\hat{\theta}(Y) \right] &\geq \frac{\left[\frac{\partial}{\partial \theta} \mathbf{E}_{\theta} \left\{ \hat{\theta}(Y) \right\} \right]^{2}}{I_{\theta}} \\ \text{where } I_{\theta} &\equiv \mathbf{E}_{\theta} \left\{ \left(\frac{\partial}{\partial \theta} \log p_{\theta}(Y) \right)^{2} \right\} \quad \text{Fisher's information} \\ &= -\mathbf{E}_{\theta} \left\{ \frac{\partial^{2}}{\partial \theta^{2}} \log p_{\theta}(Y) \right\} \end{aligned}$$

Information lower bound attained $\Leftrightarrow p_{\theta}(y)$ a single-parameter exponential pdf.

(k) Cramér-Rao Lower Bound (CRLB)

If the estimator is unbiased, i.e., $\mathbf{E}_{ heta}\left\{\hat{ heta}(Y)
ight\}= heta$, then information inequality becomes

$$\mathbf{Var}_{\theta}\left[\hat{\theta}(Y)\right] \ \geq \ \frac{1}{I_{\theta}}$$

Estimators which attain the CRLB are called *efficient*.

- 2. Procedure for finding a MVUE:
 - 1. Find a complete sufficient statistic T(y) for θ .
 - 2. Find $\hat{g}(y)$, any unbiased estimator for $g(\theta)$.
 - 3. Form $\tilde{g}(T(y)) = \mathbf{E}_{\theta} \{ \hat{g}(Y) | T(Y) = T(y) \}.$

if

3. Notes:

- Finding a complete sufficient statistic for a non-exponential family can be quite difficult.
- Form your unbiased estimator by exploiting the statistics of your observations. This is part *ad hoc* procedure, part magic.
- $\bullet\,$ The information inequality tells you what is the best you can do given the observations.

$$\begin{array}{l} \theta(y) \text{ achieves the CRLB then } \theta(y) = \theta_{\text{\tiny MVUE}}(y) \\ \text{but } \hat{\theta}(y) = \hat{\theta}_{\text{\tiny MVUE}}(y) \Longrightarrow \quad \text{CRLB achieved} \end{array}$$