

Maximum Likelihood Example
Estimation of Signal Amplitude

Setup: $Y_k = N_k + \mu s_k$ for $k = 1, \dots, n$
 $\underline{N} \sim \mathcal{N}(\underline{0}, \sigma^2 I)$
 s_k known

Objective: Estimate $\mu \in \mathbb{R}$ and $\sigma^2 \in (0, \infty)$.

Recall the MVUE's:

$$\hat{\mu}_{\text{MVUE}}(\underline{y}) = \frac{1}{n} \frac{\sum_{k=1}^n s_k y_k}{s^2} \quad \text{where} \quad \bar{s}^2 = \frac{1}{n} \sum_{k=1}^n s_k^2$$

$$\hat{\sigma}_{\text{MVUE}}^2(\underline{y}) = \frac{1}{n-1} \sum_{k=1}^n (y_k - \hat{\mu}_{\text{MVUE}}(\underline{y}) s_k)^2$$

$$\begin{bmatrix} T_1(\underline{y}) \\ T_2(\underline{y}) \end{bmatrix} = \begin{bmatrix} \sum_{k=1}^n s_k y_k \\ \sum_{k=1}^n y_k^2 \end{bmatrix} \quad \text{complete sufficient statistics}$$

Form likelihood equation:

$$\log p_{\theta}(\underline{y}) = -\frac{1}{2} \log(2\pi) - \frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{k=1}^n (y_k - \mu s_k)^2$$

Two equations, two unknowns:

$$\frac{\partial}{\partial \mu} \log p_{\theta}(\underline{y}) = \frac{1}{\sigma^2} \sum_{k=1}^n (y_k - \mu s_k) s_k \bigg|_{\substack{\mu = \hat{\mu}_{\text{ML}} \\ \sigma^2 = \hat{\sigma}_{\text{ML}}^2}} = 0$$

$$\frac{\partial}{\partial \sigma^2} \log p_{\theta}(\underline{y}) = -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{k=1}^n (y_k - \mu s_k)^2 \bigg|_{\substack{\mu = \hat{\mu}_{\text{ML}} \\ \sigma^2 = \hat{\sigma}_{\text{ML}}^2}} = 0$$

Solving the first equation for $\hat{\mu}_{\text{ML}}$ yields (likelihood global max)

$$\hat{\mu}_{\text{ML}}(\underline{y}) = \frac{1}{n} \frac{\sum_{k=1}^n s_k y_k}{s^2} = \hat{\mu}_{\text{MVUE}}(\underline{y})$$

Then using $\hat{\mu}_{\text{ML}}$ determine $\hat{\sigma}_{\text{ML}}^2$, the likelihood global max is

$$\begin{aligned} \hat{\sigma}_{\text{ML}}^2(\underline{y}) &= \frac{1}{n} \sum_{k=1}^n (y_k - \hat{\mu}_{\text{ML}}(\underline{y}) s_k)^2 \\ &= \frac{n-1}{n} \hat{\sigma}_{\text{MVUE}}^2(\underline{y}) \end{aligned}$$

So can see that $\hat{\sigma}_{\text{ML}}^2(\underline{y})$ is *biased*.

Performance: μ unknown, σ^2 known

Determine the estimator variance and Fisher's information...

$$\begin{aligned}\mathbf{Var}_\theta(\hat{\mu}_{\text{ML}}(\underline{Y})) &= \frac{\sigma^2}{n\bar{s}^2} \quad (\text{MVUE result previously derived}) \\ I_\theta &= \mathbf{E}_\theta \left\{ -\frac{\partial^2}{\partial \mu^2} \log p_\theta(\underline{Y}) \right\} \\ &= \mathbf{E}_\theta \left\{ -\frac{\partial^2}{\partial \mu^2} \left[-\frac{1}{2\sigma^2} \sum_{k=1}^n (Y_k - \mu s_k)^2 \right] \right\} \\ &= \frac{1}{\sigma^2} \sum_{k=1}^n s_k^2 = \frac{n\bar{s}^2}{\sigma^2} \\ \text{CRLB} &= \frac{1}{I_\theta} = \mathbf{Var}_\theta(\hat{\mu}_{\text{ML}}(\underline{Y}))\end{aligned}$$

Makes sense, since with $\theta = \mu$ we can write,

$$\frac{\partial}{\partial \theta} \log p_\theta(\underline{y}) = k(\theta) [\hat{\theta}_{\text{ML}}(\underline{y}) - \theta] \quad \text{with} \quad k(\theta) = I_\theta$$

meaning CRLB achievable.

Performance: μ known, σ^2 unknown

Determine the estimator variance and Fisher's information...

$$\begin{aligned}\hat{\sigma}_{\text{ML}}^2(\underline{y}) &= \frac{1}{n} \sum_{k=1}^n (y_k - \mu s_k)^2 \\ \mathbf{E}_\theta \{ \hat{\sigma}_{\text{ML}}^2(\underline{Y}) \} &= \sigma^2 \quad \text{unbiased} \\ \mathbf{Var}_\theta \{ \hat{\sigma}_{\text{ML}}^2(\underline{Y}) \} &= \frac{2\sigma^4}{n} \\ I_\theta &= \mathbf{E}_\theta \left\{ -\frac{\partial^2}{(\partial \sigma^2)^2} \log p_\theta(\underline{Y}) \right\} \\ &= \mathbf{E}_\theta \left\{ -\frac{n}{2\sigma^4} + \frac{1}{\sigma^6} \sum_{k=1}^n (Y_k - \mu s_k)^2 \right\} \\ &= -\frac{n}{2\sigma^4} + \frac{n\sigma^2}{\sigma^6} = \frac{n}{2\sigma^4} \\ \Rightarrow \text{CRLB} &= \frac{1}{I_\theta} = \mathbf{Var}_\theta(\hat{\sigma}_{\text{ML}}^2(\underline{Y}))\end{aligned}$$

Again makes sense, since with $\theta = \sigma^2$ we can write,

$$\frac{\partial}{\partial \theta} \log p_\theta(\underline{y}) = k(\theta) [\hat{\theta}_{\text{ML}}(\underline{y}) - \theta] \quad \text{with} \quad k(\theta) = I_\theta$$

meaning CRLB achievable.

Performance: μ and σ^2 both unknown

Determine the estimator variances and Fisher's information...

$$\begin{aligned}\mathbf{Var}_\theta \{ \hat{\sigma}_{ML}^2(\underline{Y}) \} &= \mathbf{Var}_\theta \left(\frac{n-1}{n} \hat{\sigma}_{MVUE}^2(\underline{Y}) \right) \\ &= \frac{(n-1)^2}{n^2} \mathbf{Var}_\theta \{ \hat{\sigma}_{MVUE}^2(\underline{Y}) \} \\ &= \frac{(n-1)^2}{n^2} \frac{2\sigma^4}{n-1} = \frac{2(n-1)\sigma^4}{n^2} \\ &< \frac{2\sigma^4}{n-1} = \mathbf{Var}_\theta \{ \hat{\sigma}_{MVUE}^2(\underline{Y}) \}\end{aligned}$$

How can ML have lower variance than MVUE???

Forcing unbiasedness increased the variance!

Consider mean-squared error:

$$\begin{aligned}\mathbf{MSE}_{ML} &= \mathbf{Var}_\theta \{ \hat{\sigma}_{ML}^2(\underline{Y}) \} + (\mathbf{E}_\theta \{ \hat{\sigma}_{ML}^2 \} - \sigma^2)^2 \\ &= \frac{\sigma^4(2n-1)}{n^2} \\ \mathbf{MSE}_{MVUE} &= \mathbf{Var}_\theta \{ \hat{\sigma}_{MVUE}^2(\underline{Y}) \} = \frac{2\sigma^4}{n-1} \\ \rightarrow \frac{\mathbf{MSE}_{MVUE}}{\mathbf{MSE}_{ML}} &= \left(\frac{n}{n-1} \right) \left(\frac{2n}{2n-1} \right) > 1\end{aligned}$$

Thus ML outperforms MVUE in both variance and MSE!

Summary

- $\hat{\mu}_{ML} = \hat{\mu}_{MVUE}$
- for μ known, $\hat{\sigma}_{ML}^2 = \hat{\sigma}_{MVUE}^2$
- with μ unknown, $\hat{\sigma}_{ML}^2$ outperformed $\hat{\sigma}_{MVUE}^2$ in variance and MSE.
- $\mathbf{Var}_\theta \{ \hat{\sigma}_{MVUE}^2 \}_{\mu \text{ known}} < \mathbf{Var}_\theta \{ \hat{\sigma}_{MVUE}^2 \}_{\mu \text{ unknown}}$