

③ We want that $\forall \varepsilon > 0, P(|X_n - \mu| > \varepsilon) \rightarrow 0$.

From Chebyshev inequality, $P(|X_n - \mu| > \varepsilon) \leq \frac{\text{var}(X_n)}{\varepsilon^2} = \frac{\sigma_n^2}{\varepsilon^2}$

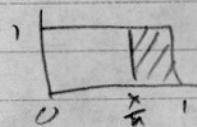
Since $\sigma_n^2 \rightarrow 0$, for any $\varepsilon > 0$: $\frac{\sigma_n^2}{\varepsilon^2} \rightarrow 0$. Thus $P(|X_n - \mu| > \varepsilon) \rightarrow 0$.
 $\forall \varepsilon > 0$

④ a) When $\omega = 1 \in \mathcal{R}$, can see that $X_n(\omega) = a_n \forall n$.
Since $a_n \rightarrow \infty$, $X_n(1) \rightarrow \infty$ and we claim $X_n \not\xrightarrow{p} 0$

b) When $\omega \in A \triangleq \{\omega \in \mathcal{R} : \omega \neq 1\}$, we claim
 $\forall \varepsilon > 0 \exists N$ s.t. $\forall n > N, |X_n(\omega)| < \varepsilon$
since choosing $N = \lceil \frac{1}{1-\omega} \rceil$ makes $X_n(\omega) = 0 \forall n > N$.
Noting $P(A) = 1$, we find that $X_n(\omega) \rightarrow 0$ for a
probability-one set of $\omega \in \mathcal{R}$. Thus $X_n \xrightarrow{as} 0$

c) Here, since $\mathcal{R} = [0, 1)$, we can find finite $N = \lceil \frac{1}{1-\omega} \rceil$
for any $\omega \in \mathcal{R}$ for which $X_n(\omega) = 0 \forall n > N$.
Thus $X_n \xrightarrow{p} 0$.

$$\begin{aligned} \textcircled{5} F_{nY_n}(x) &= P[nY_n \leq x] = 1 - P[nY_n > x] = 1 - P[Y_n > \frac{x}{n}] \\ &= 1 - P[\min(X_1, \dots, X_n) > \frac{x}{n}] = 1 - P[X_1 > \frac{x}{n}, X_2 > \frac{x}{n}, \dots, X_n > \frac{x}{n}] \\ &= 1 - P[X_1 > \frac{x}{n}]^n \quad \text{since independent} \\ &= \begin{cases} 1 - (1 - x/n)^n, & 0 \leq \frac{x}{n} < 1 \\ 1, & 1 \leq \frac{x}{n} \\ 0, & \frac{x}{n} \geq 0 \end{cases} \end{aligned}$$



Since $\lim_{n \rightarrow \infty} 1 - (1 - \frac{x}{n})^n = 1 - e^{-x}$ and $\lim_{n \rightarrow \infty} \{x : 0 \leq \frac{x}{n} < 1\} = \mathbb{R}^+$

we find

$$F_{nY_n}(x) \rightarrow \begin{cases} 1 - e^{-x}, & x \geq 0 \\ 0, & x < 0 \end{cases} \triangleq F_Y(x)$$

Thus $nY_n \xrightarrow{d} Y$, where Y is given by above F_Y